

Belief Propagation – recap

• general message equations:

$$\mu_{C \to i}(X_i) = \sum_{X_C \setminus X_i} \psi_C(X_C) \prod_{j \in C, j \neq i} \mu_{j \to C}(X_j) ,$$
$$\mu_{i \to C}(X_i) = \prod_{D \in \nu(i), D \neq C} \mu_{D \to i}(X_i)$$

beliefs:

$$b_C(X_C) := \psi_C(X_C) \prod_{i \in C} \mu_{i \to C}(X_i) , \quad b_i(X_i) := \prod_{C \in \nu(i)} \mu_{C \to i}(X_i)$$

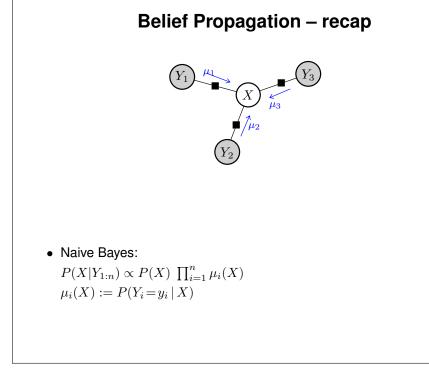
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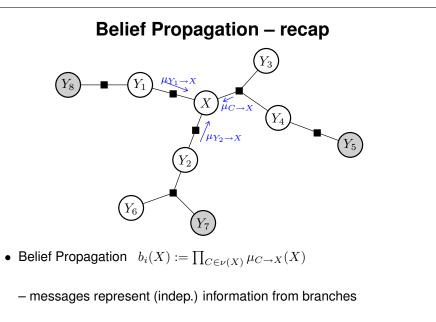
• special case: pair-wise factors \rightarrow variable-to-variable messages:

$$\mu_{j \to i}(X_i) = \sum_{X_j} \psi_C(X_i, X_j) \prod_{k: k \neq i} \mu_{k \to j}(X_j)$$

• special case: separators \rightarrow factor-to-factor messages:

$$\mu_{D\to C}(X_i) = \sum_{X_D \setminus X_i} \psi_D(X_D) \prod_{E: E \neq C} \mu_{E \to D}(X_{E \cap D}) ,$$





- messages are the temporary terms $t_k(X)$ that arise when eliminating (Elim.Alg.) a branch (\rightarrow exactness on trees) 4/30

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Hidden Markov Models

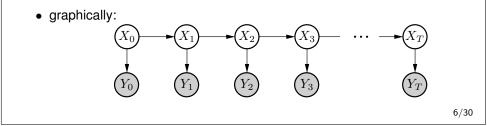
used in

- speech recognition
- molecular biology sequences
- linguistic sequences (e.g. part-of-speech tagging)
- multi-electrode spike-train analysis
- tracking objects through time
- ... motivate with data ...

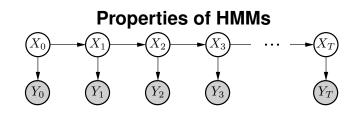
Hidden Markov Models

- HMM definition:
 - a (temporal) sequence of random variables $X_0, ..., X_T$, each with the same domain dom (X_t)
 - to each X_t and observation RV associated $Y_i,$ each with same domain $\mathrm{dom}(Y_t)$
 - the joint distribution

$$P(X_0, ..., X_T, Y_0, ..., Y_T) = P(X_0) \cdot \prod_{t=1}^T P(X_t | X_{t-1}) \cdot \prod_{t=0}^T P(Y_t | X_t) .$$



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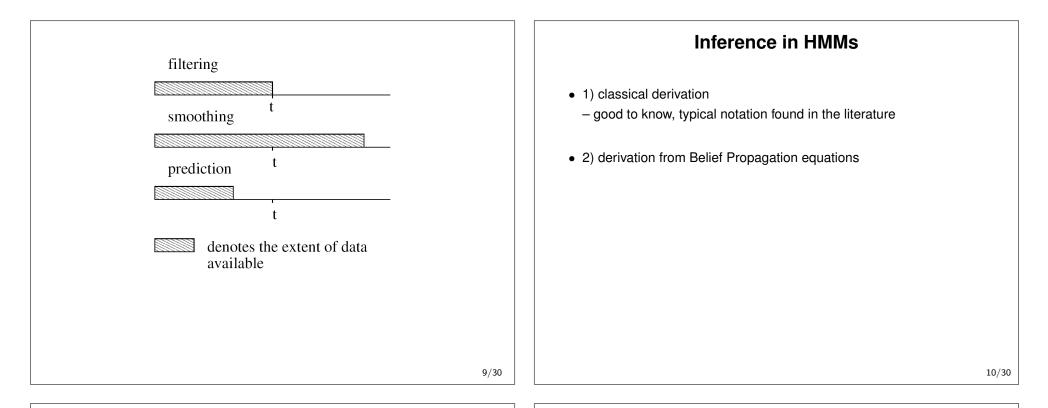
- Markov property:
 - for all $a \leq t$ and b > t
- $-X_a$ conditionally independent from X_b given X_t
- Y_a conditionally independent from Y_b given X_t

- the future is independent of the past given the present

- note that conditioning on Y_t does not yield any conditional independences

different inference problems in HMMs

- $P(x_{0:T} | y_{0:T})$ inferring hidden state given $y_{0:T}$
- $P(x_t | y_{0:T})$ marginal of above
- $P(x_t | y_{0:t})$ filtering
- $P(x_t | y_{0:a}), t > a$ prediction
- $P(x_t | y_{0:b}), t < b$ smoothing
- $P(y_{0:T})$ likelihood calculation
- Find sequence $x_{0:T}^*$ that maximizes $P(x_{0:T} | y_{0:T})$ [Viterbi alignment]



Inference in HMMs

• classical derivation

$$P(x_t | y_{0:T}) = \frac{P(y_{0:T} | x_t) P(x_t)}{P(y_{0:T})}$$

= $\frac{P(y_{0:t} | x_t) P(y_{t+1:T} | x_t) P(x_t)}{P(y_{0:T})}$
= $\frac{P(y_{0:t}, x_t) P(y_{t+1:T} | x_t)}{P(y_{0:T})}$
= $\frac{\alpha(x_t) \beta(x_t)}{P(y_{0:T})} =: \gamma(x_t)$

$$\begin{aligned} \alpha(x_t) &= P(y_{0:t}, x_t) = \phi(y_t) \ P(y_{0:t-1}, x_t) \ , \quad \phi(x_t) \equiv P(y_t | x_t) \\ &= \phi(y_t) \sum_{x_{t-1}} P(x_t | x_{t-1}) \ \alpha(x_{t-1}) \\ \beta(x_t) &= P(y_{t+1:T} | x_t) = \sum_{x+1} P(y_{t+1:T} | x_{t+1}) \ P(x_{t+1} | x_t) \\ &= \sum_{x+1} \left[\beta(x_{t+1}) \ \phi(y_{t+1}) \right] \ P(x_{t+1} | x_t) \end{aligned}$$

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Inference in HMMs

• derivation from BP equations – variable-to-variable message:

$$\mu_{j \to i}(X_i) = \sum_{X_j} \psi_C(X_i, X_j) \prod_{k: k \neq i} \mu_{k \to j}(X_j) ,$$

• message in the HMM case

$$\mu_{t-1 \to t}(x_t) = \sum_{x_{t-1}} P(x_t | x_{t-1}) \ \mu_{t-2 \to t-1}(x_{t-1}) \ \phi(x_{t-1})$$
$$\mu_{t+1 \to t}(x_t) = \sum_{x_{t+1}} P(x_{t+1} | x_t) \ \mu_{t+2 \to t+1}(x_{t+1}) \ \phi(x_{t+1})$$
$$b(x_t) = \mu_{t-1 \to t}(x_t) \ \phi(x_t) \ \mu_{t+1 \to t}(x_t)$$

belief = product of message from past, future, and cur. observation!

• compare to classical:

$$\alpha(x_t) \equiv \mu_{t-1 \to t}(x_t) \ \phi(x_t) \quad \Rightarrow \quad \mu_{t-1 \to t}(x_t) \equiv P(y_{0:t-1}, x_t)$$
$$\beta(x_t) \equiv \mu_{t+1 \to t}(x_t) \equiv P(y_{t+1:T} \mid x_t)$$

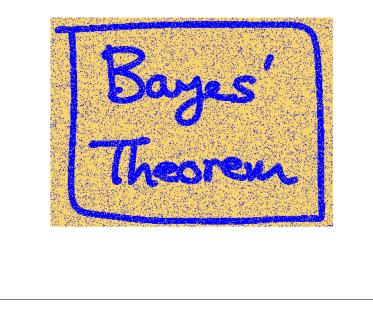
(asymmetry w.r.t. $|x_t|$ stems from asymmetry of the factors $P(x_t|x_{t-1}))_{12/30}$

HMMs

• ... demo on binary data ...

Markov Random Fields

• image denoising example:



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Markov Random Fields

• assume every pixel is a binary (black/white) random variable Let $I = \{0, ..., W\} \times \{0, ..., H\}$ be be the index set (height×width) we have binary random variables X_i for all $i \in I$ representing the pixels of the true image

we have binary randon variable Y_i representing the observations (camera snapshot)

• assume neighboring pixels are coupled

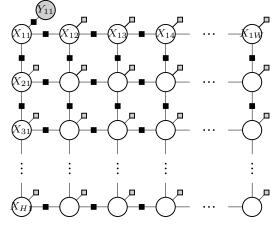
$$P(x_I, y_I) \propto \prod_{(ij)} \psi(x_i, x_j) \cdot \prod_{i \in I} \phi(x_i, y_i)$$

with couplings

$$\psi(x_i, x_j) = \begin{cases} \varrho & x_i = x_j \\ 1 - \varrho & \text{else} \end{cases} \phi(x_i, y_i) = \begin{cases} \epsilon & x_i = y_i \\ 1 - \epsilon & \text{else} \end{cases}$$

Markov Random Fields

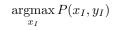
as a factor graph

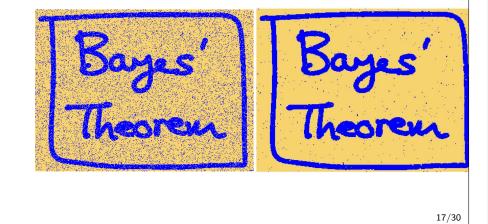


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Markov Random Fields

- image denoising is an inference problem
 - for given camera image y_I compute the most probable true image

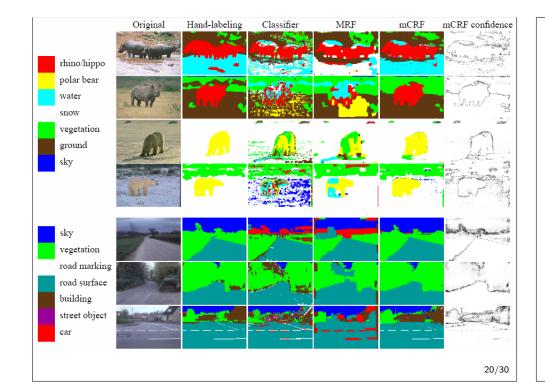




other applications

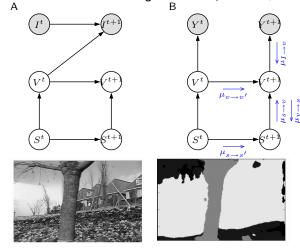
- google "conditional random field image"
 - Multiscale Conditional Random Fields for Image Labeling (CVPR 2004)
 - Scale-Invariant Contour Completion Using Conditional Random Fields (ICCV 2005)
 - Conditional Random Fields for Object Recognition (NIPS 2004)
 - Image Modeling using Tree Structured Conditional Random Fields (IJCAI 2007)
 - A Conditional Random Field Model for Video Super-resolution (ICPR 2006)

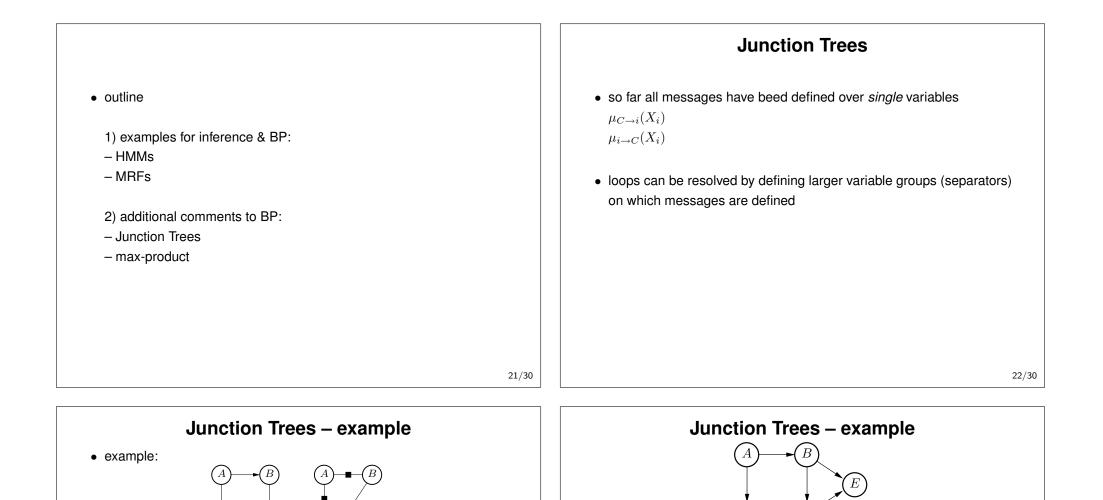
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other applications

inference for motion segmentation (Toussaint, Willert, BMVC 2007)
 A
 B





• joint variable *B* and *C* to a single "separator"



- mathematically: a variable substitution: rename the tuple (B, C) as a single random variable $\psi_1(A, B, C) = P(B|A) P(A) P(C|A)$ $\psi_2(B, C, D) = P(D|B, C)$ this still reprents the same old joint distribution P(A, B, C, D) - only factored in a different way 23/30 a variable can be contained in multiple separators – but only along a *running intersection*

Junction Tree Algorithm

- Algorithm to automatically find separators and coupling factors (=junctions) to form a tree
- graph theoretical formulation:
 - moralize a Bayes Net (= form the factor graph)
- triangulate the graph (= insert additional links/combine variables to separators)
- generate tree of maximal cliques (maximal spanning tree algorithm)
- here: use Elimination Algorithm to find the Junction Tree

Elimination Algorithm

$$P(x_{1}, x_{6})$$

$$= \sum_{x_{2}} \sum_{x_{3}} \sum_{x_{4}} \sum_{x_{5}} \psi(x_{1}, x_{2}) \psi(x_{3}, x_{1}) \psi(x_{2}, x_{4}) \psi(x_{3}, x_{5}) \psi(x_{2}, x_{5}, x_{6})$$

$$= \sum_{x_{2}} \sum_{x_{3}} \sum_{x_{4}} \psi(x_{1}, x_{2}) \psi(x_{3}, x_{1}) \psi(x_{2}, x_{4}) \sum_{x_{5}} \psi(x_{3}, x_{5}) \psi(x_{2}, x_{5}, x_{6})$$

$$= \sum_{x_{2}} \sum_{x_{3}} \sum_{x_{4}} \psi(x_{1}, x_{2}) \psi(x_{3}, x_{1}) \psi(x_{2}, x_{4}) t_{1}(x_{2}, x_{3}, x_{6})$$

$$= \sum_{x_{2}} \sum_{x_{3}} \psi(x_{1}, x_{2}) \psi(x_{3}, x_{1}) t_{1}(x_{2}, x_{3}, x_{6}) \sum_{x_{4}} \psi(x_{2}, x_{4})$$

$$= \sum_{x_{2}} \sum_{x_{3}} \psi(x_{1}, x_{2}) \psi(x_{3}, x_{1}) t_{1}(x_{2}, x_{3}, x_{6}) t_{2}(x_{2})$$

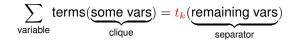
$$= \sum_{x_{2}} \psi(x_{1}, x_{2}) t_{2}(x_{2}) \sum_{x_{3}} \psi(x_{3}, x_{1}) t_{1}(x_{2}, x_{3}, x_{6})$$

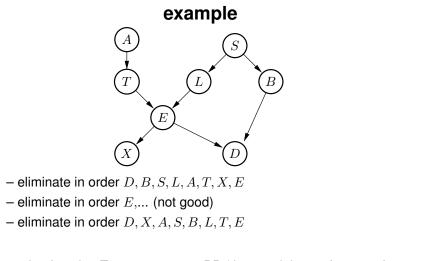
$$= t_{4}(x_{1}, x_{6})$$

Elimination \rightarrow Junction Tree

Elimination Algorithm:

- · determine an elimination order
- "simulate" the iterative process of
 - eliminating a variable
 - adding a new temporary factor $t_k(...)$ to the factor list
- keep track of the terms!





 on the Junction Tree, we can use BP (the special case factor-to-factor message equations) to do exact inference.

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comments

• naming conventions

here	other places
exact BP on trees	sum-product algorithm, message passing algorithm, inward-outward
loopy BP	BP

• finding max configurations of random variables:

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\operatorname*{argmax}_{x_{1:n}} P(X_{1:n} = x_{1:n})
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- max-product algorithm: replace \sum by \max in the message equations!
- (numerical stability: transfer to log scale and replace \prod by $\sum,$ max-sum algorithm)
- read Bishop's chapter 8 (course webpage)

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summary

- so far:
 - Bayes Nets & Factor Graphs
 - inference (Elimination, Belief Propagation)
- next big topic:
 - learning!